



LUND UNIVERSITY

Time and space order effects in timed brightness discrimination of paired visual stimuli

Patching, Geoffrey; Englund, Mats; Hellström, Åke

Published in:

Proceedings of the 24th Annual Meeting of the International Society for Psychophysics

2008

[Link to publication](#)

Citation for published version (APA):

Patching, G., Englund, M., & Hellström, Å. (2008). Time and space order effects in timed brightness discrimination of paired visual stimuli. In *Proceedings of the 24th Annual Meeting of the International Society for Psychophysics* International Society for Psychophysics.

Total number of authors:

3

General rights

Unless other specific re-use rights are stated the following general rights apply:

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
- You may not further distribute the material or use it for any profit-making activity or commercial gain
- You may freely distribute the URL identifying the publication in the public portal

Read more about Creative commons licenses: <https://creativecommons.org/licenses/>

Take down policy

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

LUND UNIVERSITY

PO Box 117
221 00 Lund
+46 46-222 00 00

TIME- AND SPACE-ORDER EFFECTS IN TIMED BRIGHTNESS DISCRIMINATION OF PAIRED VISUAL STIMULI

Geoffrey R. Patching, Mats P. Englund and Åke Hellström.

Department of Psychology, Stockholm University, SE-106 91 Stockholm, Sweden.

E-mail: grp@psychology.su.se

Abstract

Despite the considerable import of both response probability and response time for testing models of choice there is a dearth of chronometric studies of time- and space-order effects in discrimination of paired visual stimuli. In this study, systematic asymmetries in discriminating the brightness of paired visual stimuli are examined by way of binary response probability scaled in terms of log-odds ratios, as well as by signed response speed (i.e., the inverse of response time with the sign of the judged difference). For two stimuli separated by a time interval, psychometric and chronometric results revealed equivalent time-order effects, and when presented simultaneously, separated by a spatial interval, no effects of space order were found. Implications of these findings for random walk and diffusion models of sensory discrimination are discussed.

Fechner (1801-1887) was among the first to discover that when two stimuli are presented for comparison people systematically overestimate the magnitude of one stimulus and underestimate the magnitude of the other. The term, time-order effect (TOE) is used to refer to such asymmetries in paired comparisons of stimuli separated by a time interval, and the term space-order effect (SOE) to asymmetries in comparisons of paired stimuli separated spatially. By convention, a positive effect is taken to refer to an overestimation of the first (or left) stimulus as compared to the second (or right) stimulus and a negative effect as an underestimation of the first (or left) as compared to second (or right) stimulus.

In brightness discrimination, the TOE has been found to change from positive to negative with increasing inter-stimulus interval (ISI) from 1-9 seconds (Meada, 1959). Concerning the SOE, Kellogg (1931) reports a negative asymmetry in split-disk brightness discrimination, in that participants tended to choose the right-half more frequently than the left despite equally balanced brightness difference. Yet, in darkness discrimination of paired luminance gradients, participants tend to choose gradients with the darkest end on the left as compared right (Mattingley et al., 1994). So, there is some preliminary evidence to suggest that asymmetries in brightness discrimination are perceptual, but precisely what processes underlie them have yet to be fully determined.

Asymmetries in sensory discrimination are all too often dismissed as bias, which may appear as a result of prejudiced decision criteria (Allen, 1977), or verbal categorization of stimuli toward the mean of the stimulus series (John, 1975). Others envisage some kind of retention loss (Link, 1978) such that the activation inspired by one stimulus is compared to a lower fidelity mental representation of the other. In regard to the SOE, similar appeal has been made by reference to known functional asymmetries in neural anatomy (Mattingley et al., 1994), noncentral fixation and scanning effects (Masin, & Agostini, 1991).

In the paired comparison of stimuli, bias is associated with the notion of an additive effect. For instance, Davidson & Beaver, (1977) extended the classic BTL model (Bradley & Terry, 1952; Luce, 1959) to predict the probability, $p(A > B | A, B)$ of choosing A over B, given that A was presented first (or left), by inclusion of a constant order effect w , which is additive

in terms of, $\ln\psi$: $\logit [p(A>B|A,B)]=\ln\psi(A)-\ln\psi(B)-\ln w$, where $\psi(\bullet)$ is a ratio scale.

Alternatively, on the basis that the TOE and SOE are perceptual phenomena, Hellström (1979) proposed a sensation weighting (SW) model that posits a weighting of activation inspired by each stimulus event and a reference level (ReL) based on generic information. Formally, this model is defined as

$$d=k\{[s_1\psi_1 + (1 - s_1)\psi_{r1}] - [s_2\psi_2 + (1 - s_2)\psi_{r2}]\}, \quad (1)$$

where d is the subjective difference between the compared stimuli, k is a scale constant, s_1 and s_2 are weighting coefficients (1 and 2 indicate the temporal order or spatial position of the stimuli; i.e., left=1, right=2), ψ_1 and ψ_2 are the sensation magnitudes of the stimuli and ψ_{r1} and ψ_{r2} are the sensation magnitudes of the ReLs.

Further clues concerning the underlying causes of the TOE and SOE and those underlying perceptual discriminations, in general, may be obtained by examination of the time taken to make the discrimination (cf. Jamieson & Petrusic, 1975). Random walk and diffusion models propose some form of sequential sampling mechanism to explain patterns of response times and response probability in timed discrimination tasks (Link, 1975, 1992; Ratcliff, 1978). According to these models the process of comparison consists of the accumulation of noisy information about the difference between stimulus values over time, until either of two boundaries (A or $-A$) is reached. Discrimination time is defined as the time from the start of the process until one boundary is reached and response probability determined by the likelihood of crossing either boundary.

Bias is assumed to arise as a result of changes in the initial state of evidence, bound rate change, or drift rate change (Diederich & Busemeyer, 2006). For instance, Link (1978) assumes that in paired comparison tasks the activation inspired by one stimulus is compared to a lower (or higher) fidelity mental referent of the other (Link, 1978, 1992). In support of this view Link (1978) re-examined Kellogg's (1931) data in terms of the predicted relationship between response probability and response time,

$$ERT_i = \frac{A}{m} [(2P_{Ai}-1) / (S_i-S_r)] + K, \quad (2)$$

where ERT_i is expected response time, A/m is an unknown slope, K is mean non-decision time, S_i is stimulus magnitude and S_r the value of the mental referent defined in units of the stimulus. The probability P_{Ai} of the random walk exceeding one boundary (A) given stimulus S_i is given by the equation of the logistic distribution function where estimates of the logistic variable can be obtained by least squares fit to S_i of the informational value $\theta_i A$, which is the logit of P_{Ai} : $\ln(P_{Ai} / (1-P_{Ai}))$.

Link (1978) found a close correspondence between ERT (Equation 2) and mean RT as reported by Kellogg (1931). However, Kellogg (1931) did not manipulate the spatial separation between stimuli nor the absolute physical magnitude of stimuli. His data show a constant asymmetry while the TOE and SOE are characterised by changes in magnitude and direction with changes in the configuration and absolute physical magnitude of the stimuli. Consequently, without due care, differences between current sequential sampling models and human ability will continue to exist because the processes identified are grounded on an oversimplification of the actual perceptual processes engaged.

Current random walk and drift diffusion models are based on the outcome of a simple comparator which merely accumulates a signed difference by subtraction. On these grounds, Experiments 1 and 2 were conducted to examine further the TOE and SOE in timed brightness discrimination of paired visual stimuli. If the comparison process is based on the simple subtraction of stimulus information, and if systematic asymmetries in brightness discrimination change in magnitude and direction with changes in the configuration and

absolute physical magnitude of the stimuli, then it is behoved of theories based on notions of bias to explain such effects. If, however, a more complex process of comparison is admitted time- and space-order effects may be explained by perceptual weighting and by inclusion of the influence of generic information.

Experimental method

Participants. 40 participants took part in Experiment 1, and 40 different participants from the same student population took part in Experiment 2. All claimed to be right handed.

Stimuli. The visual stimuli were defined as circular spots with constant 5 mm diameter, presented on a gamma corrected 21" video monitor with a background illumination held constant at 0.01 cd/m². In Experiment 1, the paired stimuli were presented successively. Each stimulus was presented at 9 different brightness levels from 3.5 to 5.9 cd/m² in 8 steps of 0.3 cd/m², for 200 msec. The ISI was 400, 800, 1600 or 3200 msec. In Experiment 2, paired stimuli were presented simultaneously, for 200 msec, separated by a spatial interval. The spatial separation between stimuli was 10, 20, 40 or 80 mm, and the brightness of each stimulus ranged from 1.5 to 7.9 cd/m² in 8 steps of 0.8 cd/m².

Design. The 9 separate brightness levels of each visual stimulus were combined semi-factorially to form a 'diamond' shape about mean physical brightness, to create 25 different stimulus pairs. In Experiment 1, the physical differences (i.e., first minus second) between the stimuli in each pair ranged from -1.2 to +1.2 cd/m² in 4 steps of 0.6 cd/m² across one diagonal, and mean stimulus brightness ranged from 4.1 to 5.3 cd/m² in 4 steps of 0.3 cd/m² across the opposing diagonal. In Experiment 2, the physical difference between the stimuli ranged from -3.2 to 3.2 in 4 steps of 1.6 cd/m² across one diagonal and mean stimulus brightness ranged from 3.1 to 6.3 in 4 steps of 0.8 cd/m² across the opposing diagonal.

In each experiment participants took part in four 30-min sessions, one on each of 4 days. Within each session, each stimulus was shown in pseudo-randomly constructed cycles of 100 items. The first 25 trials (Exp.1) and 100 trials (Exp. 2) were designated as practice trials, immediately after which each participant was required to complete a further 400 trials per session. New random orders were used for each participant and each session.

In Experiment 1, 20 participants were instructed to press a left response key if they perceived the first stimulus to be brighter than the second, or a right key if they perceived the second stimulus to be brighter than the first. The other 20 participants were instructed to use the reverse stimulus-response assignment. In Experiment 2, all participants were instructed to respond by pressing the left response key if they perceived the left stimulus to be brighter than the right, and the right key if they perceived the right stimulus to be brighter than the left.

Procedure. At the beginning of each session participants were presented with written instructions on the video monitor. The importance of responding as quickly and as accurately as possible was stressed. The inter-trial interval was 3000 msec. On the average, participants took 30 minutes to complete each session.

Results

In Link's analysis (Equation 2) the drift rate μ is estimated by $m(S_i - S_r)$, where m is a constant. The actual perceived stimulus difference is not measured directly but represented by $\theta_i A$, which needs only to be monotonically related to μ . Therefore, ignoring the non-decision part of RT, an alternative way of thinking about μ is in terms of $\mu^* = A/RT$ for first brightest responses and $\mu^* = -A/RT$ for second brightest responses; termed, signed response speed (SRS).

Table 1. Mean log-odds ratios of the proportion first (left) brightest responses over second (right) brightest responses and mean signed response speed for the 5 brightness levels by ISI.

Experiment 1: Successively presented stimuli									
Mean brightness (cd/m ²)	ISI (msec)	Mean log-odds ratio				Mean signed response speed			
		400	800	1600	3200	400	800	1600	3200
4.1		-0.06	0.20	0.15	0.11	-0.52	1.27	0.70	0.59
4.4		-0.06	0.16	0.09	0.01	-0.62	1.01	0.53	-0.06
4.7		-0.14	0.06	0.06	-0.02	-0.87	0.34	0.53	-0.22
5.0		-0.21	-0.01	-0.12	-0.16	-1.06	0.13	-0.44	-0.91
5.3		-0.23	-0.06	-0.08	-0.27	-1.01	-0.01	-0.25	-1.3
Experiment 2: Simultaneously presented stimuli									
Mean brightness (cd/m ²)	ISI (mm)	10	20	40	80	10	20	40	80
3.1		-0.07	0.03	0.07	0.12	-0.63	0.31	0.79	1.24
3.9		-0.16	0.01	0.10	0.13	-1.55	0.01	0.10	1.06
4.7		-0.07	0.02	0.14	0.13	-0.63	0.07	1.55	1.23
5.5		-0.13	-0.08	0.09	0.21	-1.12	-0.76	1.13	1.87
6.3		-0.04	-0.03	.015	0.21	-0.29	-0.20	1.44	1.72

Here, SRS is calculated as 100/RT for first brightest response, and -100/RT for a second brightest response. The use of SRS reduces the influence of outliers and can be scaled in a similar, directly comparable, manner as response probability.

To investigate the nature of the relationship between $\text{logit}(p)$ and *SRS*, stepwise polynomial 3rd-degree regressions were fit to each participant's data over each stimulus condition and ISI (Exp.1), and spatial separation (Exp. 2), using mean *SRS* as the independent variable. On the group average, all cubic and quadratic terms in both Experiments 1 and 2 proved to be statistically nonsignificant, indicating that *SRS* and $\text{logit}(p)$ are linearly related.

Now, consider the affect of changes in average stimulus brightness and ISI on the direction and magnitude of the TOE and SOE, as shown in Table 1. For successively presented stimuli, asymmetries can be seen to change systematically in magnitude and direction with changes in stimulus brightness and ISI, both in terms of log-odds ratios and mean SRS. For simultaneously presented stimuli, there is some indication that asymmetries go from negative to positive with increased spatial separation between the stimuli but no indication of any systematic variation with changes in average stimulus brightness.

To examine changes in the magnitude and direction of the TOE and SOE the data were analysed in terms of Hellström's SW model (Equation 1). On the basis of the SW model s_1-s_2 indicates the direction of asymmetry, and the relation $s_1 \neq s_2$ was tested by (two tailed) one sample t-tests. The results of this analysis are shown in Table 2.

Further, analysis of the influence of response assignment on the magnitude and direction of s_1-s_2 revealed no statistically reliable effects (all $ps>.05$).

Table 2. Weights s_1-s_2 estimated by logistic regression of binary responses and std regression of SRS on log physical brightness (cd/m²) of the first (left) and second (right) stimuli.

Experiment 1: Successively presented stimuli									
ISI(msec)	Response frequency				SRS				
	400	800	1600	3200	400	800	1600	3200	
s_1-s_2	-0.74	-1.25	-1.26	-1.61	-0.02	-0.05	-0.05	-0.08	
P value for $(s_1-s_2) \neq 0$.02	.001	.001	.001	n.s.	.003	.012	.001	
Experiment 2: Simultaneously presented stimuli									
ISI(mm)	10	20	40	80	10	20	40	80	
s_1-s_2	0.16	-0.11	0.03	0.09	0.004	-0.009	0.009	0.009	
P value for $(s_1-s_2) \neq 0$	n.s	n.s	n.s	n.s	n.s.	n.s.	n.s.	n.s.	

Table 3: Psychometric function of stimulus difference and linear fit of Equation 2

Experiment 1: Successively presented stimuli								
ISI (msec)	400		800		1600		3200	
	$0=\theta A=$	$(2P_{Ai}-1)$	$0=\theta A=$	$(2P_{Ai}-1)$	$0=\theta A=$	$(2P_{Ai}-1)$	$0=\theta A=$	$(2P_{Ai}-1)$
Least squares fit	X Δ s+C	μ	X Δ s+C	μ	X Δ s+C	μ	X Δ s+C	μ
S_r	.14	.38	-.09	-.06	-.06	.06	.05	-.29
Multiple R (SSE)	.94 (.42)	.88 (1227)	.91 (.45)	.88 (898)	.94 (.18)	.89 (704)	.94 (.20)	.89 (633)
Slope (SE)	.93 (.15)	-294 (99)	.92 (.16)	24 (61)	.77 (.12)	57 (31)	.67 (.11)	-135 (110)
Intercept (SE)	-.20 (.12)	880 (99)	.04 (.13)	834 (25)	.01 (.09)	820 (15)	-.07 (.10)	862 (39)
Experiment 2: Simultaneously presented stimuli								
Spatial sep. (mm)	10		20		40		80	
S_r	.21	.53	.006	.43	-.29	-.09	-.38	-.58
Multiple R (SSE)	.98 (1.14)	.86 (2565)	.98 (.53)	.86 (2083)	.98 (.21)	.86 (1338)	.96 (.23)	.89 (713)
Slope (SE)	.73 (.08)	280 (115)	.73 (.07)	278 (93)	.54 (.04)	120 (74)	.36 (.05)	263 (175)
Intercept (SE)	-.21 (.18)	554 (33)	-.05 (.14)	526 (32)	.15 (.11)	569 (21)	.20 (.11)	623 (29)

Additional tests of the extended BTL and SW models were conducted by regression of binary responses and separately SRS on $[\ln(\Phi_1)-\ln(\Phi_2)]$ and $[\ln(\Phi_1)+\ln(\Phi_2)]$, where Φ_1 and Φ_2 are the physical brightness values of the first (left) and second (right) stimuli in cd/m^2 . This analysis revealed significant coefficients for $[\ln(\Phi_1)-\ln(\Phi_2)]$, all $ps < .05$, and for Experiment 1 statistically significant and marginally significant coefficients obtained also for $[\ln(\Phi_1)+\ln(\Phi_2)]$, implying $s_1 \neq s_2$. Specifically, for stimuli separated by an ISI of 3200 msec, $t(388)=1.59$, $p=.056$, $\beta=1.92$ (response probability) and $t(388)=1.62$, $p=.053$, $\beta=0.11$ (SRS), and ISI of 1600 msec, $t(391)=1.51$, $p=.066$, $\beta=1.82$ (response probability), $t(391)=1.51$, $p=.066$, $\beta=0.11$ (SRS), which suggests that time-order effects in brightness discrimination do not arise merely as result of an additive bias (cf. Davidson & Beaver, 1977).

Wave theory. Following the procedures developed by Link (1978) an average estimate of each participant's mental referent, S_r , was obtained by solving the least squares fit at which $\theta_i A = 0$, as a function of the brightness difference (Δs) in cd/m^2 between the stimuli in each pair giving 80 data points for each condition per participant. Equation 2 was then fit to mean RT as a function of Δs . As shown in Table 3, this analysis revealed a good least squares fit to S_i of the informational value $\theta_i A = \ln(P_{Ai} / (1-P_{Ai}))$, and reasonable correspondence between ERT (Equation 2), and mean RT for each condition as obtained empirically in Experiments 1 and 2.

Diffusion model analysis is currently ongoing using the DMAT toolbox (Vandkerckhove & Tuerlinckx, 2008). In the first instance, six nested models of increasing complexity were fit to the data as a function of stimulus brightness difference in cd/m^2 ; namely, i) all parameters fixed, ii) linear regression of Δs on drift rates by ISI, start position fixed midway between the two boundaries iii) linear regression of Δs on drift rates by ISI, start position free to vary, iv) logistic regression of Δs on drift rates by ISI, start position free to vary, and inter-trial range of start position free to vary, v) drift rates free to vary, start position free to vary, inter-trial range of start position free to vary, and vi) all seven parameters free to vary. The results of this analysis show no significant improvement in fit beyond linear regression of stimulus difference on drift rates allowing for changes in the intercept and slope with changes in ISI and variation in the start position of the diffusion process.

Conclusion

The findings of Experiments 1 and 2 show characteristic time- but not space-order effects in timed brightness discrimination of visual stimuli. The TOE was found to change in magnitude

and direction with changes in the temporal separation between stimuli and with changes in the absolute physical magnitude of the stimuli, whereas statistical analysis of the SOE revealed no systematic changes in magnitude and direction with changes in the absolute physical magnitude of the stimuli. Taken together, these data provide further support for the view that the TOE is a perceptual phenomenon not explainable in terms of simple response bias, verbal categorization of stimulus values toward the mean, or by the idea that the activation inspired by one stimulus is compared to a lower (or higher) fidelity mental referent of the other. The use of SRS permits consideration of current random walk and diffusion models without assumptions concerning the simple subtraction of stimulus information and hence realistic modification of these models to allow for explanations of the paired comparison of stimuli based on sensation weighting and influence of generic information.

References

- Allen, L.G. (1977). The time-order error in judgments of duration. *Canadian Journal of Psychology*, 31, 24-31.
- Bradley, R.A., & Terry, M.E. (1952). Rank analysis of incomplete block designs: I. The method of paired comparisons. *Biometrika*, 39, 324-345.
- Davidson, R.R., & Beaver, R.J. (1977). On extending the Bradley-Terry model to incorporate within-pair order effects. *Biometrics*, 33, 693-702.
- Diederich, A., & Busemeyer, J.R., (2006). Modeling the effects of payoff on response bias in a perceptual discrimination task: Bound-change, drift-rate-change or two-stage-processing hypothesis. *Perception & Psychophysics*, 68, 194-207.
- Hellström, Å (1979). Time errors and differential sensation weighting. *Journal of Experimental Psychology: Human Perception and Performance*, 5, 460-477.
- Hellström, Å (1985). The time-order error and its relatives: Mirrors of cognitive processes in comparing. *Psychological Bulletin*, 97, 35-61.
- Hellström, Å (2003). Comparison is not just subtraction: Effects of time- and space- order on subjective stimulus difference. *Perception & Psychophysics*, 65, 1161-1177.
- Jamieson D.G., & Petrusic, W.M. (1975). Presentation order effects in duration discrimination. *Perception & Psychophysics*, 17, 197-202.
- John, I.D. (1975). A common mechanism mediating the time-order error and cross-over effect in comparative judgments of loudness. *Australian Journal of Psychology*, 27, 51-60.
- Kellogg, W.M. (1931). The time of judgment in psychometric measures. *American Journal of Psychology*, 241, 1-52.
- Link, S.W. (1975). The relative judgment theory of two choice response time. *Journal of Mathematical Psychology*, 12, 114-135.
- Link, S.W. (1978). The relative judgment theory of the psychometric function. In J.Requin (Ed.), *Attention & Performance VII*. (pp. 619-630). Hillsdale, NJ: Erlbaum.
- Link, S.W. (1992). *The wave theory of difference and similarity*. Hillsdale, NJ: Erlbaum.
- Luce, R.D. (1959). *Individual choice behaviour: A theoretical analysis*. New York: Wiley.
- Maeda, H. (1959). On the inhibitory effects of extrapolated stimuli in the successive comparison of brightness (I). *Japanese Journal of Psychology*, 30, 3-20.
- Masin, S.C., & Agostini, A. (1991). Attentional scanning and space errors. *Perception & Psychophysics*, 50, 285-289.
- Mattingley, J.B., Bradshaw, J.L., Nettleton, N.C., & Bradshaw, J.A. (1994). Can task specific perceptual bias be distinguished from unilateral neglect? *Neuropsychologia*, 32, 805-817.
- Ratcliff, R. (1978). A theory of memory retrieval. *Psychological Review*, 85, 59-108.
- Vandkerckhove, J., & Tuerlinckx, F. (2008). Diffusion model analysis with MATLAB: A DMAT primer. *Behavior Research Methods*, 40, 61-72.