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Incidence of hospital norovirus outbreaks and infections using two

surveillance methods

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1

ABSTRACT

Objective: To evaluate two different methods of surveillance and to estimate the incidence of norovirus (NoV) outbreaks in hospitals.

Design: Prospective observational study.

Setting: All 194 hospital wards in southern Sweden during two winter seasons 2010-2012.

Methods: Clinical surveillance based on outbreak reports of two or more clinical cases, with symptom onset within five days, was compared with laboratory surveillance based on positive norovirus results among inpatients. At least two NoV positive patients sampled within five days at a ward defined a cluster. Outbreak reports including at least one NoV positive case and clusters including at least one NoV positive patient with five or more days from ward admission to sampling were defined as NoV outbreaks.

Results: During the study periods 135 NoV outbreaks were identified; 74 were identified by both clinical and laboratory surveillance; 18 were only identified by outbreak reports and 43 were identified by laboratory surveillance only. The outbreak incidence was 1.0 (95% CI, 0.8-1.2) and 0.5 (95% CI, 0.3-0.6) per 1000 admissions for the two different seasons respectively. To correctly identify NoV outbreaks the sensitivity and positive predictive value (PPV) of the clinical surveillance was 68 % and 88 % and 86 % and 81 % for the laboratory surveillance. **Conclusion:** The addition of laboratory surveillance significantly improves outbreak

surveillance and provides a more complete estimate of NoV outbreaks in hospitals.

Laboratory surveillance can be recommended for evaluation of clinical surveillance.

INTRODUCTION

Norovirus (NoV) is a major cause of gastroenteritis worldwide ¹⁻³ and accounting for approximately 75-90 % of all gastroenteritis outbreaks in healthcare settings. ⁴⁻⁶ Hospital NoV outbreaks cause excess morbidity among vulnerable inpatients may lead to severe consequences. ^{7,8} For healthcare facilities the outbreaks result in shortage of available beds, ill staff and economic loss. ^{6,9}

The impact of hospital NoV outbreaks has become more evident in Sweden and other industrialized countries over the last two decades, parallel with the introduction of new virus strains. Now outbreaks are a recurrent challenge to hospitals, sepecially during the cold winter season when the number of cases and outbreaks peak. Surveillance of NoV outbreaks in healthcare settings is important for early recognition and immediate infection control action and to evaluate outbreak impact and preventive measures. Despite the importance of this pathogen in healthcare settings, data on incidence of hospital NoV infections and outbreaks is still limited and the surveillance systems diverse. Surveillance has mostly been based on either reporting of clinical cases and outbreaks of gastroenteritis or laboratory reporting of NoV positive samples. Both methods have inherent limitations, as the former is dependent of compliance to report and the latter of sampling frequency and relation to outbreaks. The two different sources of data and methods to detect NoV outbreaks have not been systematically compared previously.

In this study, we estimated the incidence of NoV outbreaks in hospitals in Southern Sweden and compared the sensitivity of the two different methods of surveillance, based on either reports of clinical outbreaks of gastroenteritis or analysis of clusters of NoV positive laboratory results.

METHODS

Design

We performed a prospective observational study of outbreaks of NoV gastroenteritis during two consecutive winter seasons from 20/11/2010 to 23/04/2011 and from 26/11/2011 to 28/04/2012 at all hospitals in Region Skåne in southern Sweden. Outbreak reports were compared with the NoV laboratory results obtained from all inpatients at any hospital in the region. The study was approved by the regional ethics committee.

Hospitals and study population

Region Skåne, with 1.2 million inhabitants, have in all eight public hospitals and one small private hospital with totally 194 inpatients wards and approximately 3300 beds. The median ward size was 16 beds (interquartile range [IQR] 11-22). During the study period 184.500 hospital admissions were registered. One regional Infection Control Team (ICT), consisting of 16 nurses and three medical officers, served all the hospitals.

Clinical outbreak surveillance

According to the clinical routine the ICT was contacted by the medical staff when an outbreak at a ward was suspected. In addition, the ICT received information about positive NoV findings in inpatients directly from the laboratory. The ICT routinely contacted the wards with NoV infected patients to advice about infection control measures. During outbreaks the ICT had daily contact with the affected wards and completed an outbreak report to record epidemiological characteristics of each outbreak. The regional guideline for NoV outbreak management recommended testing at least 2-3 cases when suspecting an outbreak. No attempt was made to increase testing during the study. Definitions used for ICT clinical outbreak reports were:

A suspected NoV case was defined as a patient or healthcare worker (HCW) with diarrhea and/or vomiting (≥ 2 episodes within 24 hours) that could not be attributed to any underlying illness or medication. A <u>confirmed NoV case</u> was a suspected case with a positive NoV test by RT-PCR.

A possible outbreak was defined as two or more suspected cases, with onset within five days of each other, with suspected transmission within the ward.

A confirmed outbreak was a possible outbreak with at least one confirmed NoV case.

An outbreak was considered to have ended after a period of seven days after the last patient was reported symptom-free.¹⁵

Laboratory surveillance

All NoV diagnostic testing in Skåne was done at the Clinical Microbiology Department by RT-PCR for NoV genogroup I and II, respectively. Information about results, sampling wards and dates was obtained from the laboratory database. Dates of admission to hospitals and wards for all NoV positive patients were obtained from the patient administration database.

A <u>NoV cluster</u> was defined as at least two patients at the same ward positive for NoV of the same genogroup with sampling dates within five days. Clusters were categorized as either ward acquired, non-ward acquired or indeterminate based on the time from ward admission to NoV sampling. If at least one of the patients had been admitted to the ward ≥5 days before the date of the sampling the cluster was defined as a <u>ward acquired cluster</u>. Clusters with patients sampled 0-1 day after ward admission only were classified as non-ward acquired. Clusters were classified as indeterminate if any patient had been admitted to the ward 2-4 days before sampling and the definition of ward acquired was not met. A cluster was considered ended after a period of nine days without any NoV positive samples. This definition was set with

two extra days to allow for clinical resolution to be comparable with the clinical surveillance definition. Only the first NoV positive sample per patient and ward was used for cluster analysis. For NoV incidence only the first positive NoV test per patient and season was used. NoV infections in individual patients were classified as nosocomial, community acquired or indeterminate based if sampled ≥ 5 , ≤ 1 or 2-4 days, respectively, after admission to the hospital, not to the ward. The delay from symptom onset to sampling was validated in a random sample of 41 of the 402 patients with nosocomial infections using their medical records. The sampling delay was 0-1 days in 36 (88 %), 2-3 days in 4 (10 %) and > 4 days in 1 (2 %) of the 41 patients.

Data analysis

Wards and periods of the reported outbreaks and clusters were crosschecked for overlapping occurrences. A cluster was considered to correspond to a reported outbreak if occurring at the same ward with overlapping dates. Ward acquired clusters, without a corresponding outbreak report, were called non-reported outbreaks. Non-reported outbreaks and NoV confirmed reported outbreaks were defined as NoV outbreaks.

Data was stored and analyzed in Epi Info v 3.5.3 and Microsoft Excel 2010. In the study all wards were categorized as either psychiatric, pediatric, surgical or medical. Incidence was calculated using total number of events. Confidence intervals was calculated using the Poisson distribution. Sensitivity (separate NoV outbreaks detected / total NoV outbreaks) and positive predictive values (PPV) (separate NoV outbreaks detected / total reported outbreaks or clusters) were calculated for laboratory surveillance using the cluster definition and the clinical surveillance using the possible outbreak definition with NoV outbreak as reference. To estimate the number of outbreaks missed by both surveillance methods capture-recapture method was used. 17,18 The capture-recapture method can be used for estimates of non-detected

occurrences by evaluating the level of overlap among two incomplete and independent surveillance methods. The probability of detection in one or both methods can also estimate the probability of no detection. Non-detected outbreaks (x) were calculated by: x = bc / (a+1), where a=outbreaks detected by both methods; b=outbreaks detected by ICT reporting only and c= outbreaks detected by laboratory surveillance only (see Figure 1). The probability for a new NoV positive inpatient, at a ward without any known ongoing outbreak, to be included in a NoV outbreak was calculated by dividing NoV outbreaks by the sum of NoV outbreaks and NoV positive patients not included in any NoV outbreak.

RESULTS

Norovirus positive patients

During the study 1156 positive samples were submitted from inpatient wards representing 895 inpatients, of which 19 had positive tests at two or three wards, resulting in 915 inpatient NoV positive tests that were used for cluster analysis. Another 14 NoV positive individuals could not be verified as inpatients by hospital records. The sex distribution among all patients and Nov positive patients was the same; 54 % female and 46 % male. The incidence of NoV infection among inpatients during the two seasons in relation to age, season, ward specialty and mode of acquisition is summarized in Table 1.

Reported norovirus outbreaks

During the two winter seasons, the ICT registered 104 outbreak reports of which 92 were confirmed as NoV outbreaks (Figure 1). Of the remaining 12 possible outbreaks, nine had ≥ 2 patients tested negative for NoV and two outbreaks with no NoV tests performed. Eighty-nine of 92 outbreak reports included complete data of number of cases and comprised 817 patient and 523 HCW cases. The median number of patients and HCW cases in these outbreaks were

6 (IQR 4-11, max 57) and 4 (IQR 1-9, max 32) respectively. The median duration was 8 days (IQR 5-12, max 73) from day of onset of the first to the last cases. Ward closure was used as a control measure in 58 % of the outbreaks.

Norovirus clusters

Of the 915 NoV positive inpatients, 693 were included into 143 NoV clusters. Of these clusters, 113 (79 %) were classified as ward acquired clusters, 13 (9 %) as non-ward acquired and 17 (12 %) as indeterminate clusters (Figure 1).

Evaluation of the surveillance systems

The reported outbreaks and laboratory defined clusters were compared by wards and dates to find corresponding results (Figure 1). A total of 135 NoV outbreaks were identified. Of these, 74 were identified by both clinical and laboratory surveillance; 18 were only identified by clinical surveillance and 43 were identified by laboratory surveillance only. The laboratory surveillance identified all 74 reported NoV outbreaks with more than one positive NoV patient. The 18 reported outbreaks not identified by the laboratory surveillance had only one NoV positive patient included per outbreak. As three reported outbreaks corresponded to double clusters and one cluster corresponded to two reported outbreaks, 116 of the 143 clusters correctly corresponded to a single NoV outbreak, 27 did not correspond to any or a separate outbreak and 19 reported outbreaks were not separately identified by laboratory surveillance. The sensitivity and PPV for NoV outbreak identification was 86 % and 81 % for laboratory surveillance and 68 % and 88 % for the clinical surveillance. Using the capture-recapture method, we estimated that additional ten outbreaks would have been missed by both surveillance methods, resulting in an estimated total number of 145 NoV outbreaks. With this

estimation the adjusted sensitivity for the laboratory and the clinical surveillance was calculated to 80 % and 63 %, respectively.

Both the 92 reported and the 43 non-reported outbreaks included a median of three NoV positive patients per outbreak (IQR 2-6 and 2-4.25 respectively).

The reported and non-reported outbreaks comprised 435 and 198 NoV positive patients respectively, of which 60 % and 52 % had nosocomial and 11 % and 24 % had community acquired infections, respectively. Ten of the 43 non-reported outbreaks occurred at infectious disease wards, in which 32 (34 %) and 42 (44 %) of 95 NoV patients had nosocomial and community acquired infections, respectively. Nosocomial, but not ward acquired, infections occurred in 1/33 (3 %) of the NoV positive patients in non-ward acquired clusters and 5/67 (7 %) in indeterminate clusters.

The probability of a new NoV positive patient at an inpatient ward without known on-going outbreak to be included in a NoV outbreak was 32 % during the study periods.

Incidence of norovirus outbreaks

During the two seasons the 135 NoV outbreaks were distributed among 79 of the 194 wards (41 %) in Region Skåne. Forty-three of these wards were affected by only one outbreak but 22 wards by two and 12 wards by \geq 3 outbreaks.

The outbreak incidence was 1.0 (95% CI, 0.8-1.2) and 0.5 (95% CI, 0.3-0.6) per 1000 admissions for the two different seasons studied. The incidence of NoV outbreaks by ward specialty and season is shown in Table 2. Medical wards had significantly (p<0.05) more outbreaks than surgical wards per 1000 admissions and 100 beds. Outbreaks at pediatric wards were rare.

DISCUSSION

In this large prospective study, comprising all inpatients wards in the entire region, we show that the impact of hospital NoV outbreaks is high. Almost half of the medical wards experienced at least one outbreak during the high incidence 2010-2011 winter season and a third of the wards during the low incidence 2011-2012 winter season.

Both surveillance methods underestimated the true NoV outbreak incidence. The sensitivity of outbreak identification based on laboratory surveillance was higher than of the existing system based on active reporting.

The outbreak incidence was similar to previous reports, from which derived data show an incidence of 0.3-0.5 outbreaks per ward-year^{6,9,19} and 2.9-7.9 per 100 beds and year.²⁰ As illustrated in our results, seasonal difference and wards included can explain some of the variations between studies. We used the outbreak definitions recommended by the British Health Protection Agency, 15 and used an equivalent cluster definition. Previous studies have used similar, but not identical, cluster definitions for outbreak identification. ^{21,22} ²³ In our setting the definitions used seem adequate considering that all reported outbreaks with more than one NoV positive patients were identified as clusters. We used five or more days from admission to hospital and ward to sampling to define nosocomial infection and ward acquired clusters, respectively, as this conservative definition has been used in previous studies. 21-24 We preferred sampling date instead of symptom onset because data is readily available and may be used in future automated processing. The non-reported ward acquired clusters were similar to reported clinical outbreaks in size, but contained more patients with community acquired infections. This was mainly due to the non-reported outbreaks at infectious disease wards, indicating that these wards, responsible for the care of community acquired NoV patients, continued admitting patients during on-going ward transmission and refrained from contacting the ICT. We still consider our data to be conservative estimates of the true

incidence due to 1) the likelihood that some of the indeterminate clusters without any corresponding report or the possible non-confirmed outbreaks also were NoV outbreaks and 2) 33 of the 402 nosocomially infected patients were not included in any reported outbreak or ward acquired cluster and 3) only 8 % of the reported outbreaks contained two or three cases but 15 % had four or five cases and 77 % had more than five cases (data not shown), indicating that smaller outbreaks are less frequently reported, as is also previously described. Furthermore, the capture-recapture method, previously used in estimates of NoV outbreak burden in England, gave additional ten outbreaks missed by both surveillance methods in our analysis. Non-compliance of the ward staff to inform the ICT of suspected outbreaks and difficulties for the ICT to identify possible outbreaks by information from single NoV positive laboratory results without analytic tools might explain why ward acquired clusters were not always recognized and reported as outbreaks. Active monitoring by regular systematic ward visits, though more resource intensive, or easy electronic reporting might improve the clinical reporting system.

In our study 44 % of all the inpatients with laboratory confirmed NoV had nosocomial infections. This is less than reported from a Danish population study²⁴ and from a Dutch hospital²¹, both with the same definition as the current study, where 63 % and 52 % of the NoV positive inpatients had nosocomial infections. In a German population study, with definitions based on symptom onset, 49 % had nosocomial infections.²⁶ Apart from seasonal variation, setting and sampling indication might explain the observed differences.

This is one of the largest studies of hospital NoV outbreaks, and the first time two different surveillance methods have been directly compared. The study is based on all in-hospital wards in the entire region, served by one microbiological laboratory, minimizing selection

bias. The study was conducted during two consecutive seasons, representative of typical high and low incidence seasons. ¹² Skåne comprises more than a million inhabitants, why we believe results are generalizable to many settings.

A limitation of the study is the uncertainty of sampling delay between symptom onset to NoV sampling as it could result in misclassification of the mode of acquisition of the infections. However, the subset validation did not indicate that misclassification should be of great significance. Only patients with a NoV test from an inpatient ward were included in the study, which might result in a false low incidence of community-acquired infections and also affect the identification of clusters, if patients tested at out-patient units were later hospitalized. No comparison of timeliness of the methods was made since time of first outbreak alert was not recorded. Time lag from sampling to availability of results was in mean 1.7 days and the median time difference from when the outbreak and cluster definition was fulfilled, and an outbreak notification theoretically could have been sent, was a three days (data not shown). Laboratory surveillance using the cluster definition is thus not perfect for rapid response. The likelihood of just one NoV positive patient at a ward without any known outbreaks to become a part of an outbreak, as calculated in the present study, might be high enough for action. This "outbreak-risk" might also be used for comparison over different seasons and regions, but needs further validation before used as a quality outcome measure.

In conclusion, this study shows that the addition of laboratory surveillance to outbreak reporting significantly improves outbreak surveillance and provides a more complete estimate of the burden of NoV outbreaks in hospitals, especially when combined with admission dates. We recommend laboratory surveillance as a method for the ICT to be informed about outbreaks not reported otherwise and to evaluate clinical surveillance systems. Better methods of surveillance will improve understanding of outbreak epidemiology in healthcare settings.

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Conflict of interest: None.

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Table 1. Incidence of confirmed NoV infection for inpatients and association between age, ward specialty and mode of acquisition for two winter seasons.

	Admissio	Number of		Acquisition mode							
	ns	NoV infections (%)		Nosocomial		Indetermina	ite	Community acquired			
		No. (%)	I*	No. (%)	I*	No. (%)	I*	No. (%)	I*		
			(95% CI)		(95% CI)		(95% CI)		(95% CI)		
Age y											
< 18	15800	43 (5)	2.7 (2.0-	6 (14)	0.4 (0.2-	14 (33)	0.9 (0.5-	23 (53)	1.5 (1.0-		
			3.7)		0.8)		1.5)		2.2)		
18-65	83600	170 (19)	2.0 (1.7-	58 (34)	0.7 (0.5-	22 (13)	0.3 (0.2-	90 (53)	1.1 (0.9-		
			2.4)		0.9)		0.4)		1.3)		
>65	85 100	682 (76)	8.0 (7.4-8-	330 (48)	3.9 (3.5-	191 (28)	2.2 (1.9-	161 (24)	1.9 (1.6-		
			6)		4.3)		2.6)		2.2)		
Season											
2010-	90800	633 (71)	7.0 (6.4-	291 (46)	3.2 (2.9-	159 (25)	1.8 (1.5-	183 (29)	2.0 (1.7-		
2011			7.5)		3.6)		2.0)		2.3)		
2011-	93 700	262 (29)	2.8 (2.5-	103 (39)	1.1 (0.9-	68 (26)	0.7 (0.6-	91 (35)	1.0 (0.8-		
2012			3.2)		1.3)		0.9)		1.2)		
Ward											
specialt											
y											
Medical	94 800	703 (79)	7.4 (6.9-	308 (44)	3.2 (2.9-	173 (25)	1.8 (1.6-	222 (32)	2.3 (2.1-		
			8.0)		3.6)		2.1)		2.7)		
Surgical	69 300	133 (15)	1.9 (1.6-	66 (50)	1.0 (0.7-	34 (26)	0.5 (0.4-	33 (25)	0.5 (0.3-		
			2.3)		1.2)		0.7)		0.7)		
Psychiat	9 500	22 (2)	2.3 (1.5-	15 (68)	1.6 (1.0-	6 (27)	0.6 (0.3-	1 (5)	0.1 (0-0.7)		
ric			3.5)		2.6)		1.4)				

Pediatric	10 900	37 (4)	3.4 (2.5-	5 (14)	0.5 (0.2-	14 (38)	1.3 (0.8-	18 (49)	1.7 (1.0-
			4.7)		1.1)		2.2)		2.6)
Overall	184 500	895	4.9 (4.5-	394 (44)	2.1 (1.9-	227 (25)	1.2 (1.1-	274 (31)	1.5 (1.3-
			5.2)		2.4)		1.4)		1.7)

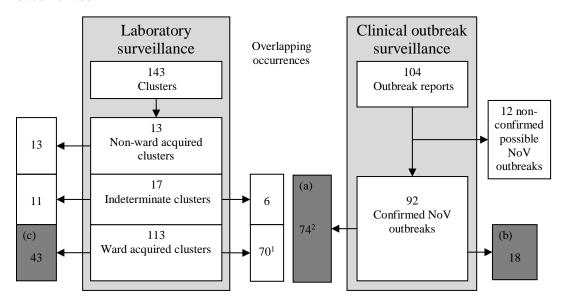
^{*} Incidence per 1000 admissions.

Table 2. Incidence of NoV outbreaks during the two seasons by main type of specialty of the wards.

Ward specialty	No. wards / beds	Admissions		No. reported / unreported outbreaks		Outbreaks* per 1000 admissions (95% CI)		Outbreaks* per 100 beds (95% CI)		Outbreaks* per ward (95% CI)		No. wards affected by outbreaks* (%	
		2010	2011	2010	2011	2010	2011	2010	2011	2010	2011	2010	2011
		2011	2012	2011	2012	2011	2012	2011	2012	2011	2012	2011	2012
Medical	90 / 1600	46 400	48 400	42 / 22	28/9	1.4 (1.1-1.8)	0.8 (0.6-1.1)	4.0 (3.1-5.1)	2.3 (1.7-3.2)	0.7 (0.6-0.9)	0.4 (0.3-0.6)	43 (47)	22 (33)
Surgical	49 / 1050	33 800	35 500	14 / 5	2 / 1	0.6 (0.4-0.9)	0.1 (0.0-0.3)	1.8 (1.2-2.8)	0.3 (0.1-0.9)	0.4 (0.2-0.6)	0.1 (0.0-0.2)	16 (28)	3 (5)
Psychiatric	40 / 500	4900	4 600	3/3	3/0	1.2 (0.6-2.7)	0.7 (0.2-2.0)	1.2 (0.5-2.7)	0.6 (0.2-1.9)	0.2 (0.1-0.3)	0.1 (0.0-0.2)	5 (14)	3 (8)
Pediatric	15 / 150	5 700	5 200	0/3	0/0	0.5 (0.2-1.6)	0	2.0 (0.6-6.2)	0	0.2 (0.1-0.6)	0	2 (13)	0
Total	194 / 3300	90 800	93 700	59 / 33	33 / 10	1.0 (0.8-1.2)	0.5 (0.3-0.6)	2.8 (2.3-3.4)	1.3 (1.0-1.8)	0.5 (0.4-0.6)	0.2 (0.2-0.3)	61 (32)	34 (18

^{*} Reported and unreported NoV outbreaks

Figure 1. Results of the two surveillance methods and the analysis of overlapping occurrences.



NoV outbreaks marked with dark grey; (a) reported confirmed outbreaks with corresponding cluster, identified by both methods, (b) reported confirmed outbreaks without corresponding cluster, identified by clinical surveillance only and (c) ward acquired cluster without corresponding outbreak report (non-reported outbreaks), identified with laboratory surveillance only.

¹One of the 70 ward acquired clusters corresponded to two reported outbreaks.

² 69 of the 74 reported outbreaks matched with ward acquired clusters and 5 matched only with indeterminate clusters. 71 of the 74 reported outbreaks matched to a single cluster but three reported outbreaks matched with two clusters.